

1 **MHC-mediated spatial distribution in brown trout (*Salmo trutta*) fry**

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29 **Abstract**

30 MHC class I-linked microsatellite data and parental assignment data for a
31 group of wild brown trout (*Salmo trutta* L.) provide evidence of closer
32 spatial aggregation among fry sharing greater numbers of MHC class I
33 alleles under natural conditions. This result confirms predictions from
34 laboratory experiments demonstrating a hierarchical preference for
35 association of fry sharing MHC alleles. Full siblings emerge from the same
36 nest (redd) and a passive kin association pattern arising from limited
37 dispersal from the nest (redd effect) would predict that all such pairs would
38 have a similar distribution. However, the present work demonstrates a
39 strong, significant trend for reduced distance between pairs of full sibling
40 fry sharing more MHC class I alleles reflecting their closer aggregation (no
41 alleles shared, $311.5 \pm (SE)21.03m$; one allele shared, $222.2 \pm 14.49m$; two
42 alleles shared, $124.9 \pm 23.88m$; $P < 0.0001$). A significant trend for closer
43 aggregation amongst fry sharing more MHC class I alleles was also
44 observed in fry pairs which were known to have different mothers and
45 were otherwise unrelated (ML-r=0) (no alleles: $457.6 \pm 3.58m$; one allele
46 ($422.4 \pm 3.86m$); two alleles ($381.7 \pm 10.72m$); $P < 0.0001$). These pairs are
47 expected to have emerged from different redds and a passive association
48 would then be unlikely. These data suggest that sharing MHC class I alleles
49 plays a role in maintaining kin association amongst full-siblings following

50 emergence. The present study demonstrates a pattern consistent with

51 MHC-mediated kin association in the wild for the first time.

52 **Keywords** MHC, brown trout, kin association,

53

54 **Introduction**

55 Kin selection theory suggests that an individual may improve its inclusive
56 fitness by behaving in a manner which enhances the reproductive success
57 of relatives (Hamilton, 1964; Keller and Ross, 1998). A prediction of the
58 theory is that the level of behavioural interactions between individual
59 organisms should be influenced by the degree of relatedness between the
60 individuals. The magnitude of the interaction can be represented by
61 Hamilton's rule; $rb-c>0$, where r is the coefficient of relatedness, b is the
62 benefit in fitness for the recipient and c is the fitness cost to the actor.
63 Following Hamilton's rule, the effectiveness of kin-mediated cooperative
64 social behaviour should be determined by the capacity of animals to
65 accurately direct such behaviour towards *bona fide* relatives.

66

67 Ancillary kin bias may occur in circumstances where interactions tend to be
68 with related rather than unrelated individuals (Sherman and Holmes,
69 1985). This may be an effect of limited dispersal from natal sites or shared
70 microhabitat preferences amongst siblings (Barnard et al., 1991; Hepper,
71 1991). This is a form of passive kin association lending itself to altruistic
72 behaviour. Other discrimination mechanisms include prior association and

73 phenotype matching (Holmes and Sherman, 1983). Kin recognition by prior
74 association refers to a requirement for a period of familiarisation between
75 the discriminator and the recipient, normally prior to weaning or fledging.
76 Phenotype matching entails the learning of phenotypic cues from self or
77 from kin that allows for kin discrimination where the target has not been
78 previously encountered. A number of studies have been conducted to
79 identify which of these two theories applied (Frederickson and Sackett,
80 1984; Grau, 1982; Halpin and Hoffman, 1987; Holmes and Sherman, 1982;
81 O'Hara and Blaustein, 1982; Waldman, 1981). The bulk of evidence
82 supports recognition by prior association (Boyd and Blaustein, 1985;
83 Dewsbury, 1982; Gavish et al., 1984; Halpin and Hoffman, 1987; Holmes
84 and Sherman, 1982; Porter et al., 1981) though there is also evidence of
85 phenotype matching (Blaustein and O'Hara, 1982; Blaustein and O'Hara,
86 1981; Grau, 1982).

87

88 The nature of cues used in kin recognition is of considerable interest to kin
89 selection. A survey of cues used for kin discrimination found that visual
90 cues were rare and that olfactory or auditory cues were more common
91 (Halpin, 1991). The reaction to these cues can be discrete, identifying that
92 the individual is related to the discriminator or not. The reaction can also
93 be proportional, in response to the degree of perceived relatedness, and

94 should be context dependent (Harris et al., 2003). In the latter case, it is
95 likely that the number of cues and the number of variants thereof is critical
96 to fine discrimination in the degree of relatedness. Hamilton (1964)
97 hypothesised the existence of a set of “recognition alleles” which entailed
98 the same or closely linked genes to both produce a phenotype that can be
99 recognised by the bearer of the same set of genes and to cause the
100 individual concerned to act altruistically to other bearers of the gene. The
101 theory was popularized by Dawkins (1976) who termed it the ‘green beard
102 effect’. A ‘green-beard’ gene produces a perceptible trait (e.g. the green
103 beard); allows recognition of other individuals with the same gene; and
104 also leads to preferential treatment towards other individuals sharing the
105 gene. The gene provides for direct recognition of copies of itself without
106 concern for average relatedness. These genes can act as ‘selfish genes’ –
107 genes which are selected for without consideration of their broader
108 adaptive value to the individual. For instance, it is feasible that otherwise
109 unrelated individuals which share ‘green-beard’ genes could recognise each
110 other as “kin” and that this would affect their behaviour towards one
111 another. These ‘green beard’ genes were first presented as hypothetical
112 mediators of altruism and kin selection. ‘Green beard’ genes have now
113 been identified in the red imported fire ant (*Solenopsis invicta*) (Keller and
114 Ross, 1998) and in yeast (*Saccharomyces cerevisiae*) (Smukalla et al., 2008).

115 The underlying gene(s) that could facilitate cues that can be used to assess
116 relatedness need to be highly variable. Attention was soon drawn to the
117 Major Histocompatibility Complex (MHC) gene family which is critical to
118 determining self from non-self in immune system function in vertebrates.
119 Classical MHC Class Ia molecules are found on the surface of all the
120 nucleated cells of the body. These are composed of a heavy and light chain
121 encoded by polymorphic MHC Class Ia genes and the invariant β 2-
122 microglobulin gene. MHC Class II genes, in contrast, are only expressed in a
123 reduced set of cells, e.g. the antigen-presenting cells such as dendritic cells,
124 B cells and macrophages. MHC Class I and II function in the presentation of
125 self and non-self peptides derived from endogenously (i.e. mutated,
126 misfolded or viral) and exogenously (e.g. bacterial or macroparasitic)
127 derived proteins, to cytotoxic T lymphocytes (CTL) or helper T cells (Th),
128 respectively.

129

130 MHC diversity is considered crucial for the ability of populations to resist
131 disease challenges (Bernatchez and Landry, 2003; Kurtz et al., 2004;
132 Muirhead, 2001; O'Brien and Evermann, 1988). MHC may also be used as
133 "recognition alleles" for helping to identify related from unrelated
134 individuals (Dawkins, 1976; Hamilton, 1964; Manning et al., 1992). This
135 process may be used in discriminatory behaviour in kin and sexual

136 selection.

137

138 The ability of organisms to detect MHC types provides a mechanism for
139 differential behaviour between MHC types. Identification of olfactory-
140 based kin recognition mediated by MHC was first demonstrated in rodents
141 (Brown et al., 1987; Yamazaki et al., 1976). Humans also appear to have the
142 capacity to discriminate amongst conspecifics based on MHC (Chaix et al.,
143 2008; Ober et al., 1997; Wedekind et al., 1995). Olfaction appears to be
144 important in behaviour in freshwater fish (Burnard et al., 2008). There has
145 also been extensive evidence for MHC-based mate selection (Agbali et al.,
146 2010; Boehm and Zufall, 2006; Consuegra and Garcia de, 2008; Eizaguirre
147 et al., 2009; Forsberg et al., 2007; Milinski et al., 2005; Milinski et al., 2010;
148 Miller et al., 2009; Neff et al., 2008; Reusch et al., 2001; Turner et al., 2009)
149 and kin discrimination (Gerlach et al., 2008; Olsen et al., 1998; Olsen et al.,
150 2002; Rajakaruna et al., 2006) in fish. However, kin discrimination has not
151 been demonstrated in the wild to date.

152

153 MHC class I (*UBA*) and class II (*DAA/DAB*) have only one expressed locus
154 each in salmonids and these loci are not linked (Aoyagi et al., 2002;
155 Grimholt et al., 2002; Shum et al., 2001; Stet et al., 2002). In this respect,
156 these genes are not a “complex”, as in other vertebrates, and are referred

157 to as “Major Histocompatibility” (MH) genes in salmonids. However, we will
158 use the more commonly used acronym “MHC” to refer to them in this
159 paper. In the laboratory setting, salmonids have been shown to prefer to
160 associate with individuals with which they share more MHC alleles (Olsen
161 et al., 1998; Olsen et al., 2002; Rajakaruna et al., 2006). We hypothesize
162 that in the wild, kin discriminatory behaviour should manifest itself as
163 spatial association based on the hierarchical sharing of MHC alleles.

164

165 The breeding behaviour of salmonids, which involves the construction of
166 nests (redds) in gravel in freshwater, suggests that young salmonids are
167 likely to be found in relatively highly related groups (Carlsson et al., 1999;
168 Hansen et al., 1997; Vera et al., 2010). This has been previously considered
169 a purely passive association due to limited dispersal from natal redds
170 (Elliot, 1987; Hansen et al., 1997) which is driven by stochastic factors
171 (Gowan et al., 1994). However, trout (*Salmo trutta* L.) were found to
172 demonstrate kin-biased distribution in both juveniles and older trout
173 (Carlsson et al., 2003). These authors suggested that the persistence of kin
174 association in older age groups implied active association although this
175 could not be proven with the data they had available to them. However,
176 Vera *et al.* (2010) found no evidence of association amongst older trout in
177 Spain.

178

179 We tested the hypothesis of active spatial association based on the
180 hierarchical sharing of MHC alleles in wild populations of trout fry,
181 separate from any passive association caused by a redd effect.

182

183 **Materials and Methods**

184 We examined spatial association of trout fry based on the sharing of MHC
185 class I-linked microsatellite marker alleles over a ~1.2km stretch of the
186 Srahrevagh River.

187 The Srahrevagh is an oligotrophic river (Dineen et al., 2007) in Co. Mayo,
188 Ireland, which has traps operated by the Marine Institute of Ireland, which
189 allow fish moving upstream and downstream to be counted and sampled
190 (Figure 1). On the 26th August 2002, 242 trout fry were obtained by
191 electrofishing the trapped stretch of the Srahrevagh River. The river was
192 divided into sections which were approximately 200m apart (range 110-
193 230m). The start and end of each section was mapped with GPS, and
194 marked with stakes. Sections were fished once in an upstream direction, by
195 one person with a backpack electrofisher, followed by one person with a
196 net. All stunned trout were netted and held in buckets. Each section was
197 fished up to a natural barrier at the top of the section (e.g. small fall) to
198 minimise fish movement out of the section. Distances between sites were

199 taken as the distance between the top of each section. Fry which were
200 found in the same site were given a distance of zero. Finer geographic
201 sampling may have improved resolution power but this was not possible in
202 the context of the current study. The distance data were tested for
203 normality using a Kolmogorv-Smirnov test. Tissue samples were stored in
204 absolute ethanol.

205

206 DNA extractions were conducted by dissecting small pieces of tissue (1-
207 5µg) from the samples and added to 0.5ml tubes containing 300µl of 10%
208 (w/v) Chelex™ solution. The mixture was heated at 99°C for 1 hour.
209 Samples were centrifuged at 3000 rpm for 3 min and then stored at -20°C.

210 All 242 individuals were screened at the selectively neutral microsatellite
211 loci *Str73* (Estoup et al., 1993), *Ssa85* and *Ssa197* (O'Reilly et al., 1996),
212 *Ssa2216* (Paterson et al., 2004), *SsOsl417* and *SsOsl85* (Slettan et al., 1995),
213 *F43* (Sanchez et al., 1996) and *Str543* (Presa and Guyomard, 1996), and a
214 microsatellite locus embedded in the untranslated 3' end of MHC class I
215 locus (*Satr-UBA*) (Grimholt et al., 2002). This dinucleotide microsatellite
216 marker is located in the 3' untranslated region (*UTR*) of the MHC class I
217 locus and has been demonstrated to be tightly linked to class I locus in
218 Atlantic salmon (*Salmo salar* L.) (Grimholt et al., 2002). It has been
219 successfully employed previously in studies in salmon (Consuegra et al.,

220 2011; de Eyto et al., 2007) and trout (Coughlan et al., 2006; Hansen et al.,
221 2007). It was not feasible to do *Satr-UBA* sequencing of the fry as *Satr-UBA*
222 has a 14,000 bp intron between exons II and III which code for the $\alpha 1$ and
223 $\alpha 1$ domains. This means that cDNA must first be derived from RNA before
224 PCR and sequencing could be conducted. The tissue samples available were
225 stored in ethanol meaning that this approach could not be undertaken.

226 *Satr-UBA* sequencing of 30 Srahrevagh trout identified 21 alleles, where
227 the linked microsatellite locus had only 13 alleles ($p < 0.0001$) (Coughlan et
228 al., 2006). The marker itself is not involved in any putative kin recognition.
229 Allele sharing at the marker is a reasonable and conservative measure of
230 allele sharing at *Satr-UBA*, as demonstrated in simulation tests below.

231

232 PCR amplifications were carried out under the following conditions in a
233 10 μ l reaction volume: 95°C 3min; (95°C for 30s, 56°C for 30s, and 72°C for
234 30s) X 30 cycles. Alleles were resolved on 18cm or 25 cm 6%
235 polyacrylamide gels, using a Li-Cor 4200 DNA sequencer. Allele sizes were
236 determined by reference to a 50-350bp size ladder and locus-specific allele
237 size standards. These allele size standards were constructed in the
238 laboratory using the full complement of allele sizes observed in pilot
239 studies, to enable consistent scoring amongst batches of individuals
240 screened for each locus. When initial genotyping was unclear due to gel

241 electrophoresis problems or weak amplification (~85 genotypes out of a
242 total of 2178), trout fry samples were re-extracted and re-screened. Large
243 allele dropout was identified as an occasional problem but large alleles
244 could usually be reliably scored after re-screening. The estimated error rate
245 was $\leq 0.5\%$ of composite genotypes per individual (Coughlan et al., 2006).

246

247 Statistical Analysis

248 A rarefaction analysis (Altmann et al., 1996; de Ruiter and Geffen, 1998)
249 was conducted using the web-based program RE-RAT (Schwacke and Rosel,
250 2005). This tested the robustness of relatedness estimates (Lynch and
251 Ritland, 1999) using the eight neutral microsatellite loci. Initially, pairwise
252 matrices of relatedness values were calculated from a single random locus.
253 After 1000 simulations, the procedure provided mean differences in
254 relatedness values with standard deviations as each additional locus was
255 added.

256

257 Each unique pair of fry was classified according to whether zero, one or
258 two MHC class I marker alleles were shared. The eight neutral
259 microsatellite markers were used to determine relatedness between fry
260 pairs. The program ML-RELATE (Kalinowski et al., 2006) was used for
261 maximum likelihood relatedness value (ML-r) estimation between pairs of

262 individuals. Full sibling pairs were identified as those showing $ML-r \geq 0.5$.
263 These full sibling pairs are expected to have emerged from the same nest
264 (redd).

265

266

267 Parental assignment data was available from sampling adults moving
268 upstream to spawn in the autumn of the previous year (2001). Parental
269 assignment data were then used to identify those pairs of fry (amongst
270 assigned fry) which did not share a mother. These fry were expected not to
271 have emerged from the same redd. Fry pairs in this sub-group with $ML-r$ of
272 0 were then identified and the remaining pairs ($ML-r > 0$) removed from the
273 analysis. These fry were then also unlikely to be paternal half siblings. This
274 was our “unrelated” group. It should be noted that the large number of
275 uncollected parents in the parental assignment data available meant very
276 few full siblings were identified, necessitating using the relatedness-based
277 approach to identify full siblings instead.

278

279 Separate analyses were then conducted for the full sibling and unrelated
280 groups. The mean distances between pairs sharing different numbers of
281 MHC alleles were calculated, Kruskal-Wallis tests and, subsequent, pairwise
282 Mann-Whitney tests between groups conducted. However, because fish

283 separated by very large distances may be equally unlikely to interact with
284 each other, we wanted to account for this through a second analysis where
285 pairs were binned by distance to “within section”, “neighbouring section”
286 (one section away) and “distant sections” (more than one section away).
287 The number of pairs sharing zero, one and two MHC alleles within each
288 distance category for the observed data were incorporated into a
289 contingency table. A bootstrap procedure was then implemented in Python
290 scripts wherein each pair of trout fry had assigned to it a random value
291 from the observed distribution of MHC marker allele shared values, with
292 replacement. The proportion of 10,000 bootstrap values which exceeded or
293 were less than the observed value in each cell of the contingency table
294 were used to determine if observed values differed significantly from
295 random expectations. The bootstrap procedure was preferable to a
296 parametric approach such as chi square test in that it provided for any
297 imbalance in the experimental design which might result in unequal
298 variances in the different groups.

299

300 Simulation study of microsatellite-UBA correlation for full sibling groups

301 We tested the validity of using the microsatellite marker as a proxy for
302 UBA allele sharing by simulating full sibling families based on the available
303 microsatellite data from this study and the *Satr-UBA* data from the same

304 river (Coughlan et al., 2006). The *Satr-UBA* data were generated from 27
305 resident brown trout adults selected to represent all microsatellite alleles
306 identified in a sample of 107 fish in 2004 from the Srahrevagh river, Co.
307 Mayo. For each microsatellite allele, the set of *UBA* alleles observed in
308 association with each microsatellite allele were compiled. The 110bp and
309 the 136bp microsatellite alleles (136bp was most common) were always
310 found with a particular *UBA* allele, *Satr-UBA*1701* and *Satr-UBA*1201*,
311 respectively.

312

313 At each iteration of the procedure, an independent full sibling data set was
314 simulated and analysed. A set of parental genotypes were then simulated
315 based on the observed MHC microsatellite data for the Srahrevagh 2002
316 trout fry sample. The numbers of males and females were independently
317 selected, a random number between 40 and 50 in both cases. Each parent
318 has a microsatellite genotype drawn at random from the observed
319 microsatellite allele frequency distribution available for the Srahrevagh
320 population. For each microsatellite allele in the parent's genotype, a
321 random *UBA* allele is drawn from the observed set of *UBA* alleles associated
322 that particular microsatellite allele (Coughlan et al., 2006). Each parent
323 then has a pair of MHC "chromosomes" with a particular combination of a
324 *UBA* and a microsatellite allele in each.

325

326 A maximum family size was selected at random between 80 and 100. We
327 then created 125 full sibling families by randomly “mating” two simulated
328 parents. The number of fry in each family was randomly selected between
329 10 and the maximum family size for each iteration. For each fry simulated,
330 a random MHC “chromosome” is inherited from the father and one from
331 the mother.

332

333 Following the generation of every full sibling family, the MHC
334 “chromosome” genotype of each fry in the family was compared with that
335 of every other fry in the family. In each pairwise comparison, the number of
336 microsatellite alleles shared amongst the pair was compared with the
337 number of *UBA* alleles shared amongst the pair. If the number is the same,
338 a “correct” score is logged under the appropriate category of microsatellite
339 alleles shared (zero, one or two). If the number of *UBA* alleles shared
340 differs, an “incorrect” score was logged under the appropriate category. At
341 the end of each iteration, the proportion of “correct” to “incorrect” scores
342 were calculated for the zero, one and two alleles shared categories. The
343 procedure was repeated across 1,000 bootstraps. These values were then
344 summarised across the 1,000 iterations of the simulation study. The
345 simulations mirrored the approach taken in our key analysis of full sibling

346 groups identified in the real data from relatedness values. The simulation
347 results gave us an idea of how well the microsatellite allele sharing patterns
348 fitted those at *UBA* within the full sibling groups simulated from the
349 Srahrevagh data. This also gave us an estimation of how well we might
350 expect the microsatellite marker to perform for identifying *UBA* allele
351 sharing in our real data.

352

353 **Results**

354

355 Distance data amongst all pairs of fry were found to be normally distributed
356 (Kolmogorov-Smirnov test, $N=29,161$, $Z=22.187$, $p = 0.000$). The rarefaction
357 analysis demonstrated that resolution power with any seven of the eight
358 neutral loci provided robust relatedness estimation. The total numbers of
359 alleles identified in the 242 fry for each locus were as follows: *Str73* (total
360 alleles: $N_A=3$, expected heterozygosity: $H_E= 0.580$), *Ssa85* ($N_A=5$, $H_E= 0.730$),
361 *Ssa197* ($N_A=8$, $H_E= 0.704$), *Ssa2216* ($N_A=6$, $H_E= 0.673$), *SsOs1417* ($N_A=21$, $H_E=$
362 0.772) and *SsOs185* ($N_A=18$, $H_E= 0.883$), *F43* ($N_A=13$, $H_E= 0.823$) and *Str543*
363 ($N_A=12$, $H_E= 0.809$). The MH class I-linked microsatellite had 11 alleles and
364 an expected Heterozygosity of 0.825.

365

366 Full sibling analysis

367 A total of 749 full sibling pairs were identified. Both individuals in each of
368 these full sibling pairs would have emerged from the same nest (redd). A
369 passive kin association pattern, arising from limited dispersal from the nest
370 (redd effect) would predict that all such pairs would have a similar
371 distribution. However, the present work demonstrates a strong, significant
372 and hierarchical trend for closer aggregation amongst full-sibling fry sharing
373 more MHC class I marker alleles, as shown by a reducing mean distance
374 between pairs sharing more alleles (no alleles shared:
375 mean=311.5±21.03m, median=170m, mean ranks=422.02, n=248; one
376 allele shared: 222.2±14.49m, median=150m, mean ranks=362.01, n=407;
377 two alleles shared: 124.9±23.88m, median = 0m, mean ranks=295.80, n=86;
378 Kruskal-Wallis test (H (df=2) =27.537, df = 2, $p < 0.0001$, Figure 2). Details of
379 all tests (including pairwise Mann Whitney U tests) can be found in Table 1.

380

381 Full sibling pairs that shared two MHC alleles were more frequent than
382 expected within a section (Bootstrap test, $p=0.0029$), not significantly
383 different from expectation at neighbouring sections ($p=0.1787$), and less
384 frequent than expected at distant sections ($p=0.0001$) (Table 2). Those
385 which shared one MHC allele were more frequent than expected within a
386 section ($p=0.0262$), but not significantly different from expectation at
387 neighbouring sections ($p=0.1699$), or distant sections ($p=0.3276$). Full

388 sibling pairs sharing no MHC alleles were significantly fewer than expected
389 within section ($p=0.0001$), not significantly different from expectation at
390 neighbouring sections ($p=0.4679$), and more frequent than expected at
391 distant sections ($p=0.0000$). Full sibling pairs which shared two MHC alleles
392 were rarely found more than two sections away from each other while full
393 sibling pairs which shared no MHC alleles were unusually common in these
394 “distant sections” (See Table 2).

395

396 Therefore, full sibling fry which emerged from the same redd not only
397 tended to be found closer to fry with which they share more MHC alleles,
398 but to have dispersed less than those sharing fewer alleles. Full siblings
399 who share two MHC alleles demonstrate a particularly marked tendency for
400 association while those that share no MHC alleles show much poorer
401 association. This suggests that sharing MHC class I alleles plays a role in
402 maintaining kin association amongst these full-siblings following
403 emergence.

404

405 Unrelated fry analysis

406 Parental assignment data was available from genotyping adults sampled
407 moving upstream to spawn in the previous year. This parental assignment
408 data was based, primarily, on complete exclusion as the only situation

409 where successful allocation of offspring is difficult to question (Jones and
410 Ardren, 2003). Offspring of the sampled males were found over a mean
411 range of ~250m while those of sampled females were found over a mean
412 range of ~400m. The “unrelated” group consisted of 15,006 pairs. We know
413 that these pairs did not emerge from the same redd as they do not share a
414 mother. Redd effect is the most likely source of passive association. A
415 paternal effect on fry spawned in adjacent sections by the same father may
416 be possible but our exclusion of fry pairs with $MI-r > 0$ should help exclude
417 half siblings.

418

419 A significant trend for closer aggregation amongst fry sharing more MHC
420 class I alleles was also observed in these unrelated pairs (zero alleles,
421 mean=457.6±3.58m, median=430m, mean ranks=7677.22; one allele,
422 mean=422.4±3.86m, median=350m, mean ranks=7341.41; two alleles,
423 mean=381.7±10.72m, median=345m, mean ranks=7198.54; Kruskal-Wallis
424 test, $H(df=2) = 25.503$, $P < 0.0001$, See Figure 3). However, no significant
425 difference was found in the pairwise Mann-Whitney test for unrelated fry
426 sharing one allele and two alleles ($U=2551245$, $z=-0.851$, $P=0.395$).

427

428 Unrelated pairs sharing two MHC alleles ($p=0.0197$) and one MHC allele
429 ($p=0.0001$) were significantly more likely to be found in the same section

430 (Table 2). Conversely, those sharing no alleles were significantly less likely to
431 be found in the same section ($p=0.0000$). No significant pattern was
432 observed for fry sharing different numbers of MHC alleles in neighbouring
433 sections ($\sim 100\text{m}$) (Table 2). Fry sharing one MHC allele were significantly
434 less likely to be found in distant sections ($\geq 200\text{m}$) ($p=0.0041$) but,
435 unexpectedly, there was no such significant pattern for fry sharing two
436 alleles ($p=0.0722$). Fry sharing no MHC alleles were significantly more
437 common in distant sections ($p=0.0009$). Overall, these data suggest that
438 MHC also significantly affects the distribution of these unrelated brown
439 trout fry. This then also appears to suggest an active association based on
440 the sharing of MHC class I alleles, given a passive association is unlikely.

441

442 Simulation study of microsatellite-*UBA* correlation for full sibling groups

443 Our simulation of full sibling groups, based on the available microsatellite
444 and *UBA* data, appeared to support the efficacy of the microsatellite
445 marker in identifying allele sharing at the *UBA* locus. Over 1,000 bootstrap
446 iterations of the simulation, full siblings which shared no microsatellite also
447 shared no *UBA* alleles 90.0% of the time (SE 0.24%), those which shared
448 one microsatellite allele shared one *UBA* allele 84.9% of the time (SE
449 0.14%) and those which shared two microsatellite alleles also shared two
450 *UBA* alleles 81.7% of the time (SE 0.27%). This implies that, when applied to

451 full sibling groups, the microsatellite marker is a good indicator of *UBA*
452 allele sharing amongst pairs of fish. The marker predicts pairs sharing two
453 alleles somewhat less accurately. If there is any bias caused by this, there is
454 no obvious reason why this would lead us to identify *UBA*-mediated kin
455 association where there is none. It seems more likely that inaccurate
456 inclusion of pairs sharing two marker alleles which do not share two *UBA*
457 alleles, for instance, would weaken the power of our analysis to identify an
458 underlying trend of association based on *UBA* allele sharing. This could
459 occur wherein pairs with no behavioural preference for one another and
460 which might be expected to be found a random distance apart, are treated
461 the same as those which share two *UBA* alleles and may demonstrate a
462 strong behavioural preference for each other.

463

464 **Discussion**

465 To our knowledge, this is the first time kin association based on the
466 sharing of MHC alleles has been demonstrated in the wild. Previous
467 laboratory studies had found evidence for this hierarchical preference for
468 fry sharing more MHC alleles (Rajakaruna et al., 2006) but that study had
469 examined MHC class II. The authors did suggest that their data implied that
470 another locus, perhaps MHC class I, was also affecting behavioural
471 preferences. Our data would appear to suggest that MHC class I is a factor

472 in kin recognition in these fish. Unrelated fry also showed signs of greater
473 association based on the sharing of MHC alleles. This could also be a
474 consequence of the use of MHC in kin recognition. Unrelated salmonids
475 were found to prefer non-kin sharing two MHC class II alleles over non-kin
476 with which they did not share alleles in the laboratory studies (Rajakaruna
477 et al., 2006). Evidence for an influence of a MHC-based kin recognition
478 system in unrelated fry, both in their behaviour and in their distribution in
479 the wild, may be indicative of a form of “green beard effect” (Dawkins,
480 1976) occurring in salmonids fish.

481

482 Laboratory studies of kin-mediated cooperative social behaviour in
483 salmonids (Brown and Brown, 1992; Brown and Brown, 1993; Brown and
484 Brown, 1996; Hedenskog et al., 2002; Olsen, 1989; Quinn and Busack, 1985;
485 Quinn and Hara, 1986) would suggest the benefits of active kin association
486 in these fry would be in the acquisition and maintenance of territories
487 following emergence from the redd. Trout which do not succeed in
488 maintaining a territory are assumed to die, with over 90% mortality during
489 this critical period (Elliot, 1994). This constitutes a powerful selective
490 pressure.

491

492 Altruistic behaviour will be most effective when directed exclusively at full

493 sibling relatives and less so when directed against half-siblings (Hamilton,
494 1964). Thus, the ability for trout fry to discriminate between full sibling
495 relatives and half-sibling relatives would be valuable. Any kin recognition
496 system will only be as effective as the resolution power achievable by the
497 genetic loci and/or phenotypic cues used. In this respect, the employment
498 of the highly polymorphic MHC loci is particularly valuable.

499

500 The resolution power of highly polymorphic loci such as MHC depends not
501 just on the number of alleles within the population at large, but on the
502 relative frequencies of each allele. The number of alleles may vary from
503 one population to the next with varying degrees of skew in the frequencies
504 of each allele. The individuals which share common alleles may identify a
505 greater proportion of conspecifics as related when they are not, with
506 fitness consequences. Conversely, individuals which encounter each other
507 and which share rare alleles may be more likely to be *bona fide* relatives.

508

509 A “rare allele advantage” for kin recognition systems has been suggested by
510 Grafen (1990). In populations which exhibit very skewed allele frequencies
511 or low polymorphism in general (perhaps due to an epizootic or population
512 bottleneck), the resolution ability of kin recognition is greatly reduced for
513 any individual possessing one or more common MHC alleles. In contrast,

514 individuals with rarer alleles enjoy a relative selective advantage arising
515 from very effective cooperative behaviours. Such an advantage is inversely
516 proportional to the frequency of the bearer's allele(s) in the population and
517 may continue to accrue over several generations. A prediction of this
518 theory would be that accurate kin recognition breaks down where MHC
519 variation is reduced in a population. Empirical evidence of kin association,
520 using a genome wide measure such as relatedness, may not be found
521 where kin recognition breaks down. However, if kin recognition is based on
522 additional genes, then kin recognition, albeit less accurate, could still
523 operate following a bottleneck event which reduced MHC variability. This
524 may provide an additional explanation for the retention of high levels of
525 polymorphism at MHC loci due to balancing selection. Heterozygosity and
526 the presence of rare MHC alleles in offspring genotypes may be beneficial
527 to survival even in the absence of direct disease challenge.

528

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536 **Captions**

537

538 Figure 1 - Map of the Srahrevagh River in Mayo, Ireland. Numbers indicate
539 the top of each section from which fish were sampled.

540 Figure 2 –Box plot of full sibling fry distribution data. Outliers (values that
541 are between 1.5 and 3 times the interquartile range) are represented by
542 circles while extreme outliers (values that are over 3 times the
543 interquartile range) are marked with asterisks. The whiskers represent the
544 minimum and maximum values which are not outliers.

545 Figure 3 - Box plot of unrelated fry distribution data.

546 Table 1 –The results of Kruskal-Wallis tests and pairwise Mann Whitney
547 tests on distance amongst pairs of fry sharing different numbers of MHC
548 alleles, for the full sibling and unrelated analyses.

549 Table 2 – Results of the bootstrap tests on the respective contingency
550 tables for the full sibling and unrelated fry

551

552

553

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554

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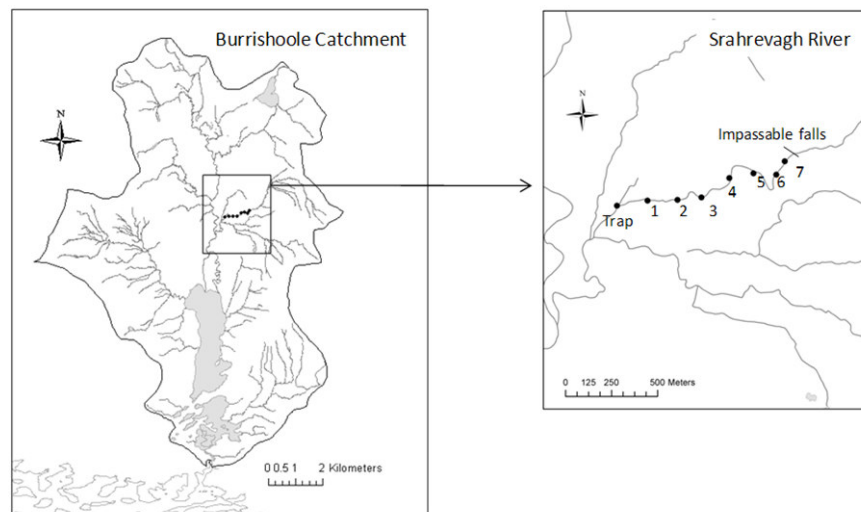
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Figure 1

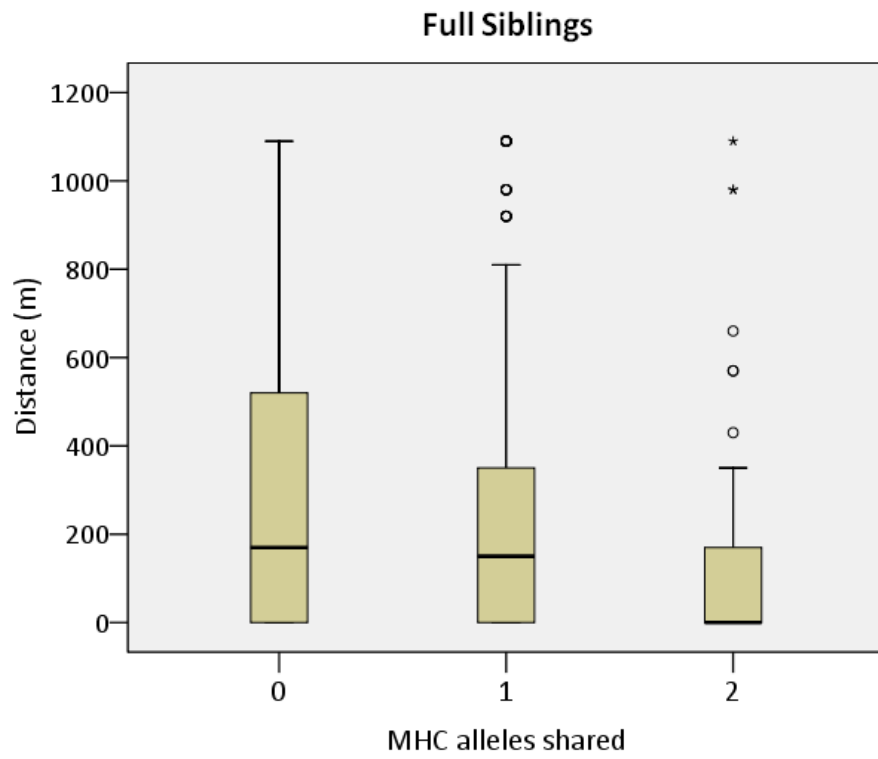


Figure 2

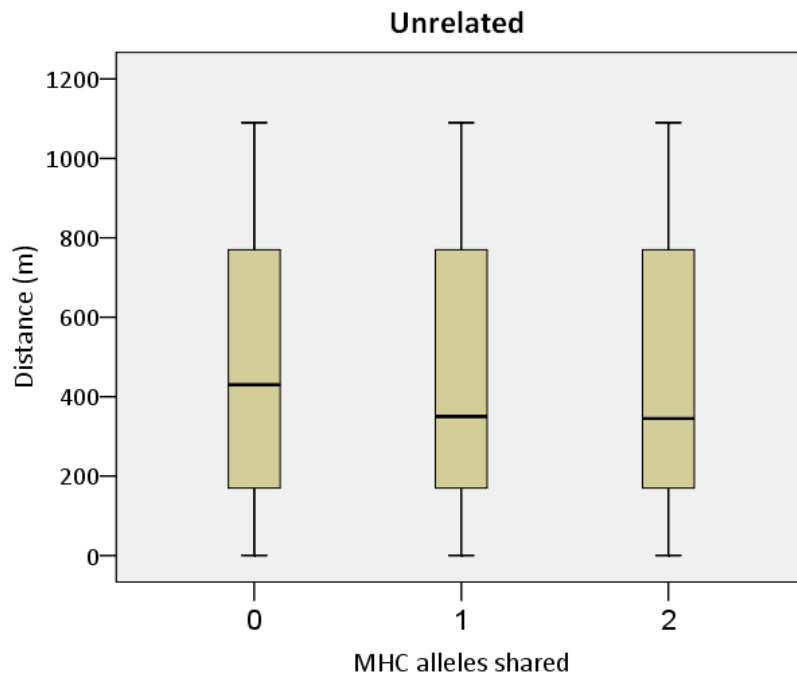


Figure 3

Table 1

MHC alleles shared		MHC alleles shared	Full Sibling			Unrelated		
			U	Z	p	U	Z	p
0	vs	1	43848.5	-3.606	0.001	24,060,000	-4.619	0.000
0	vs	2	7289	-4.896	0.000	2,771,857	-2.970	0.004
1	vs	2	14408.5	-2.714	0.006	2551245	-0.851	0.395
			Kruskal-Wallis p<0.0001			Kruskal-Wallis p<0.0001		

Table 2

Bootstraps = 10,000	MHC alleles shared	Full Sibling			Unrelated		
		Observed (proportion)	Bootstraps less than observed	p	Observed (proportion)	Bootstraps less than observed	p
Distant sections (≥200m away)	0	100 (0.435)	10000	0.0000	4,923 (0.521)	9,991	0.0009
	1	120 (0.522)	3276	0.3276	4,059 (0.430)	41	0.0041
	2	10 (0.043)	1	0.0001	461 (0.049)	722	0.0722
Neighbouring section (~100m)	0	66 (0.355)	5321	0.4679	1,686 (0.492)	570	0.0570
	1	94 (0.505)	1699	0.1699	1,555 (0.453)	8,886	0.1114
	2	26 (0.140)	8213	0.1787	189 (0.055)	8,002	0.1998
Within section	0	82 (0.252)	1	0.0001	967 (0.453)	0	0.0000
	1	193 (0.594)	9738	0.0262	1034 (0.485)	9,999	0.0001
	2	50 (0.154)	9971	0.0029	132 (0.062)	9,803	0.0197